

Exact asymptotic results for the Bernoulli matching model of sequence alignment

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Finding analytically the statistics of the longest common subsequence (LCS) of a pair of random sequences drawn from c alphabets is a challenging problem in computational evolutionary biology. We present exact asymptotic results for the distribution of the LCS in a simpler, yet nontrivial, variant of the original model called the Bernoulli matching (BM) model. We show that in the BM model, for all c , the distribution of the asymptotic length of the LCS, suitably scaled, is identical to the Tracy-Widom distribution of the largest eigenvalue of a random matrix whose entries are drawn from a Gaussian unitary ensemble.

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Sequence alignment is one of the most useful quantitative methods used in evolutionary molecular biology [1–3]. The goal of an alignment algorithm is to search for similarities in patterns in different sequences. A classic and much studied alignment problem is the so-called “longest common subsequence” (LCS) problem. The input to this problem is a pair of sequences $\alpha = \{\alpha_1, \alpha_2, \dots, \alpha_i\}$ (of length i) and $\beta = \{\beta_1, \beta_2, \dots, \beta_j\}$ (of length j). For example, α and β can be two random sequences of the four base pairs A, C, G, T of a DNA molecule, e.g., $\alpha = \{A, C, G, C, T, A, C\}$ and $\beta = \{C, T, G, A, C\}$. A subsequence of α is an ordered sublist of α (entries of which need not be consecutive in α), e.g., $\{C, G, T, C\}$, but not $\{T, G, C\}$. A common subsequence of two sequences α and β is a subsequence of both of them. For example, the subsequence $\{C, G, A, C\}$ is a common subsequence of both α and β . There can be many possible common subsequences of a pair of sequences. The aim of the LCS problem is to find the longest of such common subsequences. This problem and its variants have been widely studied in biology [4–7], computer science [8–10,2], probability theory [11–16], and more recently in statistical physics [17–19]. A particularly important application of the LCS problem is to quantify the closeness between two DNA sequences. In evolutionary biology, the genes responsible for building specific proteins evolve with time and by finding the LCS of the same gene in different species, one can learn what has been conserved in time. Also, when a new DNA molecule is sequenced *in vitro*, it is important to know whether it is really new or if it already exists. This is achieved quantitatively by measuring the LCS of the new molecule with another existing already in the database.

For a pair of fixed sequences of length i and j respectively, the length $L_{i,j}$ of their LCS is just a number. However, in the stochastic version of the LCS problem one compares two random sequences drawn from c alphabets and hence the length $L_{i,j}$ is a random variable. A major challenge over the last three decades has been to determine the statistics of $L_{i,j}$ [11–15]. For equally long sequences ($i=j=n$), it has been proved that $\langle L_{n,n} \rangle \approx \gamma_c n$ for $n \gg 1$, where the averaging is performed over all realizations of the random sequences. The constant γ_c is known as the Chvátal-Sankoff constant. There

exist several bounds [12,14,15], a conjecture due to Steele [13] that $\gamma_c = 2/(1+\sqrt{c})$ and a recent proof [16] that $\gamma_c \rightarrow 2/\sqrt{c}$ as $c \rightarrow \infty$. Unfortunately, no exact results are available for the finite-size corrections to the leading behavior of the average $\langle L_{n,n} \rangle$, for the variance, and also for the full probability distribution of $L_{n,n}$. Therefore it is important to find other variants of this LCS problem that may be analytically tractable for the full distribution of the matching length.

Computationally, the easiest way to determine the length $L_{i,j}$ of the LCS of two arbitrary sequences of lengths i and j [in polynomial time $\sim O(ij)$] is via using the recursive algorithm [2,19]

$$L_{i,j} = \max[L_{i-1,j}, L_{i,j-1}, L_{i-1,j-1} + \eta_{i,j}], \quad (1)$$

subject to the initial conditions $L_{i,0} = L_{0,j} = L_{0,0} = 0$. The variable $\eta_{i,j}$ is either 1 when the characters at the positions i (in the sequence α) and j (in the sequence β) match each other, or 0 if they do not. Note that the variables $\eta_{i,j}$'s are not independent of each other. To see this consider the simple example of the matching of two strings $\alpha = AB$ and $\beta = AA$. One has by definition: $\eta_{1,1} = \eta_{1,2} = 1$ and $\eta_{2,1} = 0$. The knowledge of these three variables is sufficient to predict that the last two letters will not match, i.e., $\eta_{2,2} = 0$. Thus, $\eta_{2,2}$ cannot take its value independently of $\eta_{1,1}$, $\eta_{1,2}$, $\eta_{2,1}$. These residual correlations between the $\eta_{i,j}$ variables make the LCS problem rather complicated. Note, however, that for two random sequences drawn from c alphabets, these correlations between the $\eta_{i,j}$ variables vanish in the $c \rightarrow \infty$ limit.

A natural question is how important are these correlations between the $\eta_{i,j}$ variables, e.g., do they affect the asymptotic statistics of $L_{i,j}$'s for large i and j ? Is the problem solvable if one ignores these correlations? These questions naturally lead to the Bernoulli matching (BM) model, which is a simpler variant of the original LCS problem where one ignores the correlations between $\eta_{i,j}$'s for all c [19]. The length $L_{i,j}^{BM}$ of the BM model satisfies the same recursion relation in Eq. (1) except that $\eta_{i,j}$'s are now independent and each drawn from the bimodal distribution: $p(\eta) = (1/c)\delta_{\eta,1} + (1 - 1/c)\delta_{\eta,0}$. This approximation is expected to be exact only in the appropriately taken $c \rightarrow \infty$ limit. Nevertheless, for fi-

nite c , the results on the BM model can serve as a useful benchmark for the original LCS model to decide if indeed the correlations between $\eta_{i,j}$'s are important or not. Unfortunately, even in the absence of correlations, the exact asymptotic distribution of $L_{i,j}^{BM}$ in the BM model has so far remained elusive, mainly due to the nonlinear nature of the recursion relation in Eq. (1). The purpose of this Rapid Communication is to present an exact asymptotic formula for the distribution of the length $L_{n,n}^{BM}$ in the BM model for all c .

So far, only the leading asymptotic behavior of the average length in the BM model is known [19] using the "cavity" method of spin-glass physics [20],

$$\langle L_{n,n}^{BM} \rangle \approx \gamma_c^{BM} n, \quad (2)$$

where $\gamma_c^{BM} = 2/(1 + \sqrt{c})$, same as the conjectured value of the Chvátal-Sankoff constant γ_c for the original LCS model. However, other properties such as the variance or the distribution of $L_{n,n}^{BM}$ remained intractable even in the BM model. Our main result is that for large n ,

$$L_{n,n}^{BM} \rightarrow \gamma_c^{BM} n + f(c)n^{1/3} \chi, \quad (3)$$

where χ is a random variable with an n -independent distribution, $\text{Prob}(\chi \leq x) = F_{\text{TW}}(x)$, which is the well-studied Tracy-Widom distribution for the largest eigenvalue of a random matrix with entries drawn from a Gaussian unitary ensemble [21]. For a detailed form of the function $F_{\text{TW}}(x)$, see [21]. We show that for all c ,

$$f(c) = \frac{c^{1/6}(\sqrt{c} - 1)^{1/3}}{\sqrt{c} + 1}. \quad (4)$$

This allows us to calculate the average including the subleading finite-size correction term and the variance of $L_{n,n}^{BM}$ for large n ,

$$\begin{aligned} \langle L_{n,n}^{BM} \rangle &\approx \gamma_c^{BM} n + \langle \chi \rangle f(c)n^{1/3}, \\ \text{Var } L_{n,n}^{BM} &\approx (\langle \chi^2 \rangle - \langle \chi \rangle^2) f^2(c)n^{2/3}, \end{aligned} \quad (5)$$

where one can use the known exact values [21], $\langle \chi \rangle = -1.7711\dots$ and $\langle \chi^2 \rangle - \langle \chi \rangle^2 = 0.8132\dots$. These exact results thus invalidate the previous attempt [19] to fit the subleading correction to the mean in the BM model with a $n^{1/2}/\ln(n)$ behavior and also to fit the scaled distribution with a Gaussian form. Note that the recursion relation in Eq. (1) can also be viewed as a $(1+1)$ -dimensional directed polymer problem [18,19] and some asymptotic results [such as the $O(n^{2/3})$ behavior of the variance of $L_{n,n}$ for large n] can be obtained using the arguments of universality [18]. However, this does not provide precise results for the full distribution that are obtained here.

It is useful to provide a synopsis of our method in deriving these results. First, we prove the results in the $c \rightarrow \infty$ limit, by using mappings to other models. To make progress for finite c , we first map the BM model exactly to a three-dimensional (3D) anisotropic directed percolation (ADP) model studied by Rajesh and Dhar [22]. This ADP model can further be mapped to a $(1+1)$ -dimensional directed polymer problem studied by Johansson [23]. For this specific directed

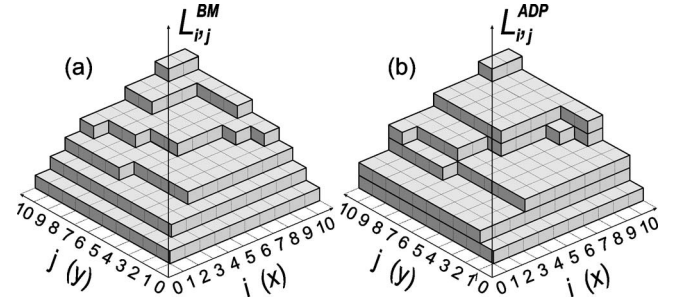


FIG. 1. Examples of (a) the BM surface $L_{i,j}^{BM} \equiv \tilde{h}(x,y)$ and (b) the ADP surface $L_{i,j}^{ADP} \equiv h(x,y)$.

polymer problem, Johansson derived exact asymptotic results for the distribution of the polymer energy. To extract the results for the BM model from those of Johansson's model, we use a simple symmetry argument which then allows us to derive our main results in Eqs. (3)–(5) for all c . As a check, we recover the $c \rightarrow \infty$ limit result obtained independently by the first method.

In the BM model, the length $L_{i,j}^{BM}$ can be interpreted as the height of a surface over the two-dimensional (2D) (i,j) plane constructed via the recursion relation in Eq. (1). A typical surface, shown in Fig. 1(a), has terracelike structures.

It is useful to consider the projection of the level lines separating the adjacent terraces whose heights differ by 1 (see Fig. 2) onto the 2D (i,j) plane. Note that, by the rule Eq. (1), these level lines never overlap each other, i.e., no two paths have any common edge. The statistical weight of such a projected 2D configuration is the product of weights associated with the vertices of the 2D plane. There are five types of possible vertices with nonzero weights as shown in Fig. 2, where $p = 1/c$ and $q = 1 - p$. Since the level lines never cross each other, the weight of the first vertex in Fig. 2 is 0.

Consider first the limit $c \rightarrow \infty$ (i.e., $p \rightarrow 0$). The weights of all allowed vertices are 1, except the ones shown by black dots in Fig. 2, whose associated weights are $p \rightarrow 0$. The number N of these black dots inside a rectangle of area $A = ij$ can be easily estimated. For large A and $p \rightarrow 0$, this number is

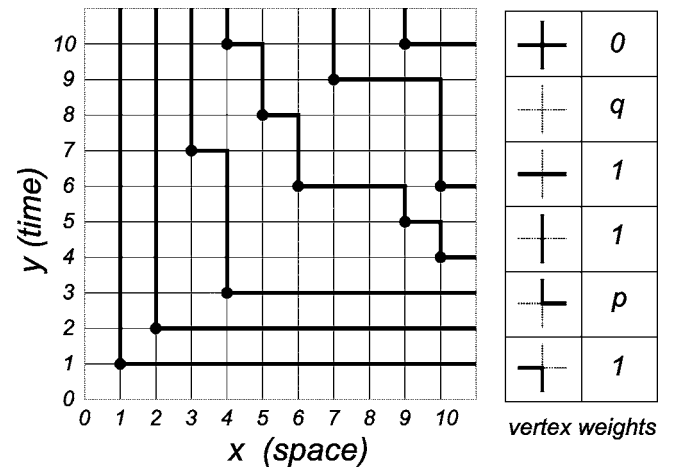


FIG. 2. Projected 2D level lines separating adjacent terraces of unit height difference in the BM surface in Fig. 1(a). The adjacent table shows the weights of all vertices on the 2D plane.

Poisson distributed with the mean $\bar{N}=pA$ (see, e.g., remark 3 in Section 4 of [24]). The Bethe ansatz analysis shows that BM corresponds to the sector of the five-vertex model [25] where the density α of empty edges in a row of vertical edges is close to the boundary $\alpha \approx 1^-$. The careful examination of the free energy near this boundary allows one to conclude that the leading contribution in p (for $p \rightarrow 0$) to \bar{N} comes exactly from the line of phase transitions in a five-vertex model. The subleading corrections to \bar{N} are of order $\sim p^{3/2}$ and are ensured by small deviations from the critical line being beyond the Poisson approximation [26].

The height $L_{i,j}^{BM}$ is just the number of level lines \mathcal{N} inside this rectangle of area $A=ij$. The problem of estimating \mathcal{N} has recently appeared in a number of interface models such as a polynuclear growth model [27] and a ballistic deposition model [28]. By using a mapping to the longest increasing subsequence (LIS) of the equally likely permutations of a set of integers and then, by applying a celebrated result due to Baik, Deift, and Johansson (BDJ) [30], it was shown [27,28] that the number of level lines \mathcal{N} inside the rectangle (for large A), appropriately scaled, has a limiting behavior, $\mathcal{N} \rightarrow 2\sqrt{\bar{N}+\bar{N}^{1/6}}\chi$, where χ is a random variable with Tracy-Widom distribution. Using $\bar{N}=pA=ij/c$, one then obtains in the limit $p \rightarrow 0$,

$$L_{i,j}^{BM} = \mathcal{N} \rightarrow \frac{2}{\sqrt{c}}\sqrt{ij} + \left(\frac{ij}{c}\right)^{1/6} \chi. \quad (6)$$

In particular, for large equal length sequences $i=j=n$, we get for $c \rightarrow \infty$

$$L_{n,n}^{BM} \rightarrow \frac{2}{\sqrt{c}}n + c^{-1/6}n^{1/3}\chi. \quad (7)$$

For finite c , while the above mapping to the LIS problem still works, the corresponding permutations of the LIS problem are not generated with equal probability and hence one can no longer use the BDJ results.

For any finite c , we can map the BM model to the ADP model [22]. In the ADP on a simple cubic lattice the bonds are occupied with probabilities p_x , p_y , and p_z along the x , y , and z axes and are all directed towards increasing coordinates. Imagine a source of fluid at the origin that spreads along the occupied directed bonds. The sites that get wet by the fluid form a 3D cluster. In the ADP problem, the bond occupation probabilities are anisotropic, $p_x=p_y=1$ (all bonds aligned along the x and y axes are occupied) and $p_z=p$. Hence, if the point (x,y,z) gets wet by the fluid then all the points (x',y',z) on the same plane with $x' \geq x$ and $y' \geq y$ also get wet. Such a wet cluster is compact and can be characterized by its bounding surface height $h(x,y)$ as shown in Fig. 1(b). It is not difficult to see that the height $h(x,y)$ satisfies the following recursion relation [22],

$$h(x,y) = \max[h(x-1,y), h(x,y-1)] + \xi_{x,y}, \quad (8)$$

where $\xi_{x,y}$'s are i.i.d. (independent and identically distributed) random variables taking non-negative integer values with $\text{Prob}(\xi_{x,y}=k) = (1-p)p^k$ for $k=0,1,2,\dots$. One can also interpret the height $h(x,y)$ in Eq. (8) as the energy of a di-

rected polymer in the x - y plane. Precisely this particular version of the polymer problem was studied by Johansson [23], who obtained the asymptotic distribution of the height for large x and y ,

$$h(x,y) \rightarrow \frac{2\sqrt{pxy+p(x+y)}}{q} + \frac{(pxy)^{1/6}}{q} \times \left[(1+p) + \sqrt{\frac{p}{xy}}(x+y) \right]^{2/3} \chi, \quad (9)$$

where $q=1-p$, and χ is a random variable with a Tracy-Widom distribution.

While the terrace-like structures of the ADP surface look similar to the BM surfaces [compare Figs. 1(a) and 1(b)], there is an important difference between the two. In the ADP model, the level lines separating two adjacent terraces can overlap with each other [22], which does not happen in the BM model. However, by making the following change of coordinates in the ADP model [22]:

$$\zeta = x + h(x,y), \quad \eta = y + h(x,y), \quad (10)$$

one gets a configuration of the surface where the level lines no longer overlap. Moreover, it is not difficult to show that the projected 2D configuration of level lines of this shifted ADP surface has exactly the same statistical weight as the projected 2D configuration of the BM surface. Denoting the BM height by $\tilde{h}(x,y) = L_{x,y}^{BM}$, one then has the identity, $\tilde{h}(\zeta, \eta) = h(x,y)$, which holds for each configuration. Using Eq. (10), one can rewrite this identity as

$$\tilde{h}(\zeta, \eta) = h[\zeta - \tilde{h}(\zeta, \eta), \eta - \tilde{h}(\zeta, \eta)]. \quad (11)$$

Thus, for any given height function $h(x,y)$ of the ADP model, one can, in principle, obtain the corresponding height function $\tilde{h}(x,y)$ for all (x,y) of the BM model by solving the nonlinear equation (11). This is, however, very difficult in practice. Fortunately, one can make progress for large (x,y) where one can replace the integer-valued discrete heights by continuous functions $h(x,y)$ and $\tilde{h}(x,y)$. Using the notation $\partial_x \equiv \partial/\partial x$ it is easy to derive from Eq. (10) the following pair of identities,

$$\partial_x h = \frac{\partial_\zeta \tilde{h}}{1 - \partial_\zeta \tilde{h} - \partial_\eta \tilde{h}}, \quad \partial_y h = \frac{\partial_\eta \tilde{h}}{1 - \partial_\zeta \tilde{h} - \partial_\eta \tilde{h}}. \quad (12)$$

In a similar way, one can show that

$$\partial_\zeta \tilde{h} = \frac{\partial_x h}{1 + \partial_x h + \partial_y h}, \quad \partial_\eta \tilde{h} = \frac{\partial_y h}{1 + \partial_x h + \partial_y h}. \quad (13)$$

We then observe that Eqs. (12) and (13) are invariant under the simultaneous transformations

$$\zeta \rightarrow -x, \quad \eta \rightarrow -y, \quad \tilde{h} \rightarrow h. \quad (14)$$

Since the height is built up by integrating the derivatives, this leads to a simple result for large ζ and η ,

$$\tilde{h}(\zeta, \eta) = h(-\zeta, -\eta). \quad (15)$$

Thus, if we know exactly the functional form of the ADP surface $h(x, y)$, then the functional form of the BM surface $\tilde{h}(x, y)$ for large x and y is simply obtained by $\tilde{h}(x, y) = h(-x, -y)$. Changing $x \rightarrow -x$ and $y \rightarrow -y$ in Johansson's expression for the ADP surface in Eq. (9), we thus arrive at our main asymptotic result for the BM model

$$L_{x,y}^{BM} = \tilde{h}(x, y) \rightarrow \frac{2\sqrt{pxy} - p(x+y)}{q} + \frac{(pxy)^{1/6}}{q} \left[(1+p) - \sqrt{\frac{p}{xy}}(x+y) \right]^{2/3} \chi, \quad (16)$$

where $p=1/c$ and $q=1-1/c$. For equal length sequences $x=y=n$, Eq. (16) then reduces to Eq. (3).

To check the consistency of our asymptotic results, we further computed the difference between the left- and the right-hand sides of Eq. (11),

$$\Delta h(\zeta, \eta) = \tilde{h}(\zeta, \eta) - h[\zeta - \tilde{h}(\zeta, \eta), \eta - \tilde{h}(\zeta, \eta)], \quad (17)$$

with the functions $h(x, y)$ and $\tilde{h}(x, y)$ given respectively by Eqs. (9) and (16). For large $\zeta = \eta$ one gets

$$\Delta h(\zeta, \zeta) \rightarrow [p^{1/3} \chi^2 / 3(1 - \sqrt{p})^{4/3}] \zeta^{-1/3}. \quad (18)$$

Thus the discrepancy falls off as a power law for large ζ , indicating that indeed our solution is asymptotically exact. We have also performed numerical simulations of the BM model using the recursion relation in Eq. (1) for $c=2, 4, 9, 16, 100$. Our preliminary results [26] for relatively small system sizes (up to $n=5000$) are consistent with our exact results in Eqs. (3)–(5).

The Tracy-Widom distribution of the random matrix theory has appeared recently in a number of problems [21,23,27–29,31]. In this Rapid Communication we have shown that it also describes the asymptotic distribution of the length in the BM model for all c . Given that the correlations in the original LCS model become negligible in the $c \rightarrow \infty$ limit, it is likely that the BM asymptotics in Eq. (7) would also hold for the original LCS model in the $c \rightarrow \infty$ limit (obtained by first taking the $n \rightarrow \infty$ limit for fixed c and then taking the $c \rightarrow \infty$ limit), though it needs to be proved rigorously. An important open problem is to determine whether the Tracy-Widom distribution also appears in the LCS problem for finite c . The precise distribution obtained here (including exact prefactors) for all c in the BM model will serve as a useful benchmark to which future simulations of the LCS problem can be compared.

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